

A genome-wide association meta-analysis identifies a novel locus at 17q11.2 associated with sporadic amyotrophic lateral sclerosis

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ABSTRACT

Identification of mutations at familial loci for amyotrophic lateral sclerosis (ALS) has provided novel insights into the aetiology of this rapidly progressing fatal neurodegenerative disease. However, genome wide association studies (GWAS) of the more common (~90%) sporadic form have been less successful with the exception of the replicated locus at 9p21.2. To identify new loci associated with disease susceptibility we have established the largest association study in ALS to date and undertaken a GWAS meta-analytical study combining 3,959 newly genotyped Italian individuals (1,982 cases, 1,977 controls) collected by SLAGEN (Italian Consortium for the Genetics of ALS) together with samples from Netherlands, USA, UK, Sweden, Belgium, France, Ireland and Italy collected by ALSGEN (the International Consortium on Amyotrophic Lateral Sclerosis Genetics). We analyzed a total of 13,225 individuals, 6,100 cases and 7,125 controls for almost 7 million single nucleotide polymorphisms (SNPs). We identified a novel locus with genome-wide significance at 17q11.2 (rs34517613 $P = 1.11 \times 10^{-8}$; OR 0.82) that was validated when combined with genotype data from a replication cohort ($P = 8.62 \times 10^{-9}$; OR 0.833) of 4,656 individuals. Furthermore, we confirmed the previously reported association at 9p21.2 (rs3849943 with $P = 7.69 \times 10^{-9}$; OR 1.16). Finally, we have estimated the contribution of common variation to heritability of sporadic ALS as ~12% using a linear mixed model accounting for all SNPs. Our results provide an insight into the genetic structure of sporadic ALS, confirming that common variation contributes to risk and that sufficiently powered studies can identify novel susceptibility loci.

INTRODUCTION

Amyotrophic lateral sclerosis (ALS) is a late onset progressive neurodegenerative disorder mainly affecting motor neurons. ALS is the most common adult onset motor neuron disease with prevalence of 5 per 100,000 and with family history, age and male gender as the major risk factors (1-4). While familial ALS is well characterized with several causative genes identified to date, the genetic architecture of the more common sporadic form is poorly understood. Previous GWAS have identified several loci associated with ALS risk such as *DPP6*, *ITPR2*, *FGGY* and *UNC13a* (5-11) that have failed to be replicated in independent populations (12-15). One exception is a locus on chromosome 9p21 (15) that has been reliably replicated (16). At this locus an expanded hexanucleotide repeat in the *C9orf72* gene has been recently identified as the causative mutation in a large proportion of familial (23-47%) as well as sporadic (~5%) ALS cases (17-20). This finding indicates that well powered studies can identify novel loci associated with ALS susceptibility. To discover further loci we designed a large GWAS meta-analysis with sufficient power to detect risk alleles with small effect sizes as observed in other neurodegenerative and complex diseases. There is evidence for a strong genetic component in sporadic ALS with heritability estimated to be 0.61 (95% CI: 0.38-0.78) in a combined study of 171 ALS twin pairs collected by three independent studies from Britain and Sweden (21), so we sought to estimate heritability derived from unrelated individuals and explained by common variation using GWAS data. Phenotype variation in complex traits is due to interactions of genetic and environmental factors, thus the quantification of the genetic variance is relevant in the study of multifactorial diseases. Heritability, the proportion of phenotype explained by genetic variance, is typically estimated in close related individuals such as twin pairs but this can inflate estimates as a consequence of epistatic interactions or shared environment (22,23). In contrast, heritability explained by associated SNPs identified in GWAS as passing accepted thresholds of significance (typically

$P \leq 5 \times 10^{-8}$) explains only a small fraction of the genetic variation in most complex diseases. The difference between the phenotypic variance explained by GWAS results and those estimated in family studies is referred to as the “missing heritability problem” and possibly is explained by incomplete linkage disequilibrium (LD) between genotyped SNPs and causal variants and/or by the presence of gene-by-gene or gene-by-environment interactions (22,23). We have estimated heritability of ALS considering all SNPs simultaneously regardless of their association with ALS phenotype and compared this with heritability from twin studies.

RESULTS

Association analyses

We analyzed genotype raw data from eight independent studies including 3,959 newly genotyped Italian individuals (1,982 cases, 1,977 controls) and 11,611 individuals (5,195 cases and 6,416 controls) from previously published studies. Full description of the sample size included in this study is reported in Table 1 and Supplementary Table S1.

Raw data for each study were assessed for quality control (QC) separately following the same criteria (see Supplementary Material and Supplementary Table S2 for full discussion).

Population structure of the individual cohorts was studied by means of principal components analysis (PCA) with EIGENSTRAT software (24,25) and outliers were identified by the projection of the first 10 principal components (Methods, Supplementary Material, Supplementary Fig. S1).

To achieve the maximal coverage we imputed genome-wide filtered originally genotyped data for each study (Methods). Overall the average number of inferred genotypes was 8,342,920 SNPs varying proportionally to the original genotyping platform. After filtering (Methods, Supplementary Material), the eight data sets included genotype data for almost 7 million SNPs

in 13,225 individuals (6,100 cases and 7,125 controls) (Supplementary Table S2). Cleaned imputed genotypes were tested for association with ALS separately using SNPTESTv2.4.0 (26) and the logistic regression analyses were adjusted by the appropriate principal component axes estimated in the individual cohorts (Supplementary Table S3). After genomic inflation control, we combined the logistic regression analysis results of each study in a meta-analysis using the program METAL (27). We adopted the “standard error” analysis scheme which combines effect size estimates (β -coefficients) across the studies weighted according to the inverse of the corresponding standard errors. As a result of the meta-analysis we observed two loci reaching genome wide threshold of significance, rs3849943 $P = 7.69 \times 10^{-9}$ (OR 1.16, 95% CI: 1.10-1.22; average posterior probability APP=0.999; MAF cases 0.268, controls 0.238) on chromosome 9p21 and SNP rs34517613 $P = 1.11 \times 10^{-8}$ (OR 0.82, 95% CI: 0.76-0.87; APP = 0.9279; MAF cases 0.108, controls 0.129) on chromosome 17q11.2.

A third locus, rs1788776, at 18q11.2 was very close to genome wide significance threshold ($P = 7.67 \times 10^{-8}$, OR 0.87, 95% CI: 0.76-0.87; APP = 0.9716; MAF cases 0.392, controls 0.362) (Table 2, Fig. 1, Fig. 2, Supplementary Fig. S2). With the exception of the locus on 9p21, loci previously found to be associated with ALS risk (5-11) did not reach genome wide significance in our meta-analysis (Supplementary Table S4) though SNP rs12608932 (*UNC13A*) at 19p13.3 had not complete coverage across datasets as it failed the QC threshold in the British data set (Supplementary Fig. S3). Additionally, examination of these previously associated loci in the Italian cohort as an independent replication study only found evidence for association for the 9p21 locus, and for the *ITPR2* locus while significant, direction of effect is the opposite to that reported in the original study (8) (Supplementary Fig S3). Only the *ITPR2* locus showed significant evidence for heterogeneity across the different cohorts of European ancestry analyzed here (Supplementary Table S5). Further larger studies are needed to confirm the role of these loci in ALS susceptibility.

As inclusion of non-confounding covariates such as gender and age can significantly reduce power to identify truly associated variants particularly when the disease prevalence is low (< 1%), we have not included non-confounding covariates in logistic regression analyses presented here (28). In confirmation of our approach, the most significant SNPs ($P \leq 1 \times 10^{-5}$) mapping in the two candidate loci at 17q and 18q showed no gender-by-SNP interaction in multi regression analysis (Methods).

For these loci we searched for secondary signals by performing logistic regression conditioning upon the most associated SNPs under an additive model. There was no evidence of independent SNPs in any of these regions (Supplementary Table S6). As the international meta-analysis included samples from Northern and Southern Europe, we tested the most associated SNPs for genetic heterogeneity quantifying the effect of the degree of variation between studies (Methods, Supplementary Material). We observed no significant heterogeneity in the distribution of allele frequencies between European populations for these SNPs (Supplementary Table S7).

As additional evidence, we tested the candidate loci at 17q and 18q in a novel independent replication cohort of 2,074 cases and 2,556 controls collected in Italy, Netherlands and Germany (see full description in the Methods, Supplementary Material and Supplementary Table S8). For each locus, we selected two variants, the most associated in the GWAS meta-analysis (imputed data) and combined joint analysis (original genotype data) (Supplementary Material, Supplementary Table S9). We found suggestive evidence for association in the independent replica analysis for SNP rs34517613 at 17q ($P = 0.055$; OR 0.89, 95% CI: 1.00-0.78) and when combined in meta-analysis with the full GWAS data the association identified in the discovery sample became more significant ($P_{\text{combined}} = 8.62 \times 10^{-9}$; OR 0.833, 95% CI: 0.891-0.778). We found no additional evidence for association at the 18q11.2 locus (Supplementary Table S10).

The locus at 17q11.2 is gene rich with a number of plausible candidates (Fig. 2). In an effort to refine associations at this locus and at 18q11.2, we searched for non synonymous SNPs in LD ($r^2 \geq 0.4$) with our two lead signals to identify possible functional variants that explained the associations with ALS risk. No non-synonymous variants were found in LD at the 17q11.2 locus; however, three SNPs, rs739439 (r^2 0.416) in the 3' untranslated region, the intronic SNPs rs35714695 (r^2 0.457) and rs34660379 (r^2 0.412) within the *SARM1* (sterile alpha and TIR-containing motif 1) gene were in LD (Table 2, Fig. 2).

At the suggestive 18q11.2 locus the lead SNP rs1788776 was only in strong LD with SNPs in the *ANKRD29* gene (Fig. 2), an ankyrin repeat domain-containing protein of undetermined function, these included missense SNPs (rs12956232, r^2 0.78; rs1788758, r^2 0.78; rs12960692, r^2 0.74; rs11662113, r^2 0.53) (Table 2, Fig. 2).

Next, we looked for cis eQTL (expression quantitative trait loci) in LD with our lead SNPs. All SNPs in LD ($r^2 \geq 0.2$) with rs34517613 and rs1788776 were analyzed for association with cis eQTL using several publicly available eQTL databases (Supplementary Material) but none reached the *P*-value threshold for significance. This should be not considered conclusive as expression data from central nervous system tissues are still rather limited across eQTL databases and further studies should be carried out.

Contribution of *C9orf72* to chromosome 9p21 association signal

As chromosome 9p21 was found to be strongly associated in our meta-analysis, we further investigated the contribution of ALS sporadic cases that also carry the expanded hexanucleotide repeat in the *C9orf72* gene to this association. Information about carriers of the pathologic expansion in *C9orf72* gene was available for 2,287 out of 6,100 (37.3%) ALS cases from the Italian, Dutch and British studies included in the meta-analysis. The frequency of the expanded repeats carriers in the sporadic cases we have analyzed progressively decreased as

you proceed from Northern to Southern Europe ranging from 8% to 4.7% respectively in the British and Italian patients (Methods, Supplementary Table S10). Our results confirm data previously reported (17-20). We performed a meta-analysis of this sample subset (2,287 sporadic cases, 4,162 controls) including and excluding cases (n=144) with the expanded hexanucleotide repeat. Significant association of SNP rs3849943 decreased from $P = 7.72 \times 10^{-5}$ to $P = 0.052$ when the carriers were excluded from the logistic regression analysis confirming that the association of 9p21 locus was largely dependent on carriers of the *C9orf72* expansion (Supplementary Fig. S4).

Heritability estimation

We used the Genome-wide Complex Trait Analysis (GCTA) method (29) to quantify additive genetic variance (narrow sense heritability) explained by all common SNPs regardless of their association with disease. GCTA uses a linear mixed-effect model that includes a random effect to estimate the polygenic component of trait variation (29). We performed GCTA analysis on 13,225 individuals (6,100 ALS cases and 7,125 controls) whose pair wise relationship was estimated to be less than 0.05 (PI_HAT) and used a restricted maximum likelihood (REML) algorithm to estimate the variance explained by imputed SNPs data. Analyses were carried out separately in the eight cohorts using a more stringent quality control threshold for imputed SNPs (MAF > 0.01, posterior probability > 0.9 and information measure > 0.8) (Methods, Supplementary Material Table S2) under the assumption of an ALS prevalence of 5 per 100,000. Heritability estimates were similar (~11.5%) for all cohorts (Supplementary Table S12) and were combined in meta-analysis modelled for random effects and weighted by sample size. The summary heritability was 0.119 (95% CI: 0.110-0.127) (Fig. 3).

Exclusion of the chromosome 9p21 locus did not affect heritability estimates. As ALS is a disease of aging and age is increasing in the human population, an increase in the prevalence of

ALS can be predicted. Consequentially the true prevalence of ALS may be higher than that used in our analysis. We have therefore calculated heritability over a range of prevalences, but for all prevalences estimation of heritability attributable to common variation does not exceed 0.25 (Supplementary Fig. S5).

DISCUSSION

In this study we present the largest GWAS meta-analysis to date in ALS. Overall we analyzed 13,225 individuals with sufficient power to capture allelic association with small effects and low MAF. The novel associated SNP rs34517613 at 17q11.2 was confirmed when genotype data from the independent replication cohort were combined in meta-analysis with full GWAS data. We have also replicated the 9p21 locus at genome wide significance; this locus was also identified by linkage studies of familial ALS patients and by analogy it may be that there is a corresponding familial locus for the 17q11.2 locus identified here, though we are not aware of any reports.

The lead SNP at this locus was in LD with three SNPs in the *SARM1* gene. Interestingly, *SARM1* orthologues, *dSarm* in *Drosophila melanogaster* and *Sarm1* in mice, have been recently found to play a direct role in an axonal self-destruction pathway, a mechanism known as Wallerian degeneration which shares morphological similarities with the axon dying-back degeneration observed in ALS and in other neurodegenerative pathologies (30,31). As axonal degeneration is an early feature of ALS progression, a mouse model for familial ALS, *SOD1G93A*, was crossed with the *Wld^S* mouse, a spontaneous mutant with phenotype of prolonged survival of injured axons due to a chromosomal rearrangement that disrupts two genes, *Ube4b* and *Nmnat1*. In the *SOD1G93A/Wld^S* the progression of axonopathies was attenuated suggesting that the Wallerian pathway could be involved in the axon loss observed in ALS (32,33).

We estimated the heritability of ALS due to common variation as ~12% based on a prevalence of 5 per 100,000 (1,2), although the true prevalence of ALS may be higher as consequence of increasing age in the human population as discussed above. We have also only considered autosomal SNPs and additional variance may be encoded on the X chromosome. Twin studies estimate heritability due to all genetic variation including rare monogenic forms of the disease whilst we have estimated heritability as a consequence of common polymorphisms either directly typed or imputed. The difference between heritability estimated from twin studies and from analysis of common SNP polygenic variation suggests a substantial role for variation not captured by genome wide association studies.

Our estimate for the heritability of ALS due to common variation is lower than that for other late-onset neurodegenerative diseases. Using similar approaches, heritability for late-onset Parkinson's disease was 0.31 (34), 0.24 for late-onset Alzheimer's disease and 0.30 for Multiple sclerosis (35). These differences in heritability due to common variants are reflected in the difficulty in identifying genome wide significant loci for ALS compared to these diseases. For example, a two stage GWA study of Parkinson's disease with a stage 1 sample size of around 6,000 individuals identified two strong association signals (36) and a study of Alzheimer's disease with a similar sized stage 1 sample identified two loci additional to the well replicated *APOE* locus (37). In comparison to these studies our combined sample size is more than 13,000 individuals suggesting that the genetic architecture of ALS may be different from other more common neurodegenerative diseases. A recent study of ALS in a Han Chinese population identified two loci of genome wide significance in a combined sample of around 5,000 (38); strikingly these loci were not replicated in populations of European ancestry, nor in the combined meta-analysis reported here. This suggests there may be considerable heterogeneity in ALS risk loci across different ethnicities, in support of this, the expanded repeat in the *C9orf72* gene has been reported to be infrequent in Asian populations (39,40).

In conclusion, we have identified a novel locus for sporadic ALS risk at 17q11.2, but not replicated the suggestive evidence for a second locus at 18q11.2 and confirmed the association at 9p21 with the expanded hexanucleotide repeat in the *C9orf72* gene. Furthermore we provide evidence from our heritability estimates that further common variation affecting ALS risk remains to be detected by current GWAS platforms and with larger cohorts but that denser genome wide assays and next generation sequencing technologies are required to detect rarer variation.

MATERIALS AND METHODS

Participating individuals

The discovery sample consisted of a novel Italian cohort collected by the Italian Consortium for the Genetics of ALS (SLAGEN) and seven independent published studies collected by ALSGEN, The International Consortium on Amyotrophic Lateral Sclerosis Genetics (5-11,15,16). All patients fulfilled the El Escorial revised criteria for ALS (41). Written informed consent according to the Declaration of Helsinki was obtained from all patients and healthy subjects participating to this study. Local ethical committees for each participating Institution approved these studies. All samples across the eight data sets were of European ancestry and a full description of demographic details is reported in the Supplementary Table S1. All patients with family history of ALS or carrying Mendelian risk genes were excluded from analyses. Screening for the expanded hexanucleotide repeat in the *C9orf72* gene were performed subsequently; mutation carries information was available for a subset of cases included in the Italian, Dutch and British cohorts (see Supplementary Material for full description).

A replication sample of 4,630 individuals (2,074 cases, 2,556 controls) was collected across Italy, the Netherlands and Germany using the same diagnostic criteria (see Supplementary Material, Supplementary Table S8).

Genotyping Procedures

Samples from the individual cohorts were genotyped on different Illumina beadchips (Illumina, CA, USA) as shown in Table 1. Additional SNPs were genotyped in the replication phase by PCR-based KASP technique by the KBioscience (UK) facility for the Italian samples (622 cases, 971 controls) or by TaqMan_2013 PCR method for the Dutch (877 cases, 1226 controls) and German samples (575 cases, 359 controls) (Supplementary Material). Cross validation between methodologies showed 100% concordance.

Statistical analysis

Quality control and meta-analysis of imputed data

Before imputation analysis quality control (QC) of samples and markers were performed separately in the eight studies using the PLINK software package (42). A detailed description of QC procedure is reported in the Supplementary Material and Supplementary Table S2.

Genotype data from British cases and controls were merged and analyzed as single cohort as fully described in the Supplementary Material and Supplementary Table S2.

Ancestry differences between individuals within each cohort were detected by principal components analysis (PCA). Principal component axes were generated by genotypes of a genome wide subset of LD independent SNPs using EIGENSTRAT software (24,25) and outliers identified by the first 10 principal components (PCs) were removed. The number of significant principal components was estimated by Tracy-Widom distribution using the program TW statistic (25) and included as covariates in the logistic regression analyses of

each study for population stratification control. Scatter plots of PC1 and PC2 showed no evidence of substructure between cases and controls within the single cohorts

(Supplementary Fig. S1). As an example, Supplementary Figure S5 shows the population structure of the novel Italian cohort by the projection of the first two PC axes.

After quality control, original genotype data of each study were tested for genomic inflation and lambda estimates resulted to be minimal ($\lambda_{(gc)} < 1.02$). Individual datasets were imputed genome wide separately using the IMPUTE.v2 program (43,44) that employs combined reference panels of known phased haplotypes provided by HapMap 3 (Feb 2009), 1,000 Genomes Project (Mar 2010) (NCBI build 36 coordinates) and the study's sample genotypes (Supplementary Material). In each cohort imputed genotypes were tested for association with ALS status by logistic regression analysis (SNPTESTv2.4.0) including the specific PCs as confounder covariates to control population stratification. For each test statistic spurious associations and genomic inflation were controlled by plotting the observed quantiles versus the expected ones in Q-Q plots and by factor $\lambda_{(gc)}$ estimate (Supplementary Fig. S7, Supplementary Table S3) (R package software).

SNP statistic data were filtered for uncertainty of inferred genotypes according to posterior probability (APP) > 0.9 and statistical information of allele frequency (info) > 0.4 (26). The average number of SNPs filtered out in each data set was 1.47 million (Supplementary Table S2). Finally, filtered SNP tables were combined in a meta-analysis and analyzed with the program METAL (27), applying the standard error scheme option that weights effect size estimates, or β -coefficients, using the inverse of the corresponding standard errors. The z -scores were calculated from the P values of each study with the sign of the z -scores indicating the direction of allelic effect in that study ($z > 0$ for odds ratio > 1). Meta-analysis was performed applying the fixed-effects model that does not account for between study heterogeneity, therefore heterogeneity was quantified by the Cochran's statistic (Q) and I^2

(METAL). P values were estimated by comparing the statistic with a X^2 distribution ($N-1$ degrees of freedom, N = number of studies) (Supplementary Material, Supplementary Table S5 and S7) (45,46). Meta-analysis was repeated stratifying by gender and testing for the gender-by-SNP interaction as described in the Supplementary Material.

For the independent replication cohort logistic regression analyses (SNPTESTv2.4.0) were performed separately in the Italian, Dutch and German samples and then combined in meta-analysis (METAL) applying standard error scheme as described above. Next, statistic tables of the three novel replication samples were combined with the discovery sample in a final meta-analysis (METAL).

Power calculation

The final meta-analysis cohort had 99.99% power to detect allelic association with an odds ratio (OR) of 1.2 and minor allele frequency (MAF) of 0.25 at genome wide significance ($P = 5 \times 10^{-8}$). For low frequency variants with MAF ranging from 0.01 to 0.04 and of small effect (OR < 1.5), there was low power to capture variants with MAF 0.01 while for SNPs with MAF 0.02 power ranged from 60% to 98% for alleles with OR of 1.3 and 1.4 respectively (Supplementary Fig. S8).

Screening of expanded repeats in *C9orf72* gene

Hexanucleotide repeat expansion data for the *C9orf72* gene were available for the Italian, Dutch and British cases (Supplementary Table S10). As previously reported carriers were defined as individuals with a range of the GGGGCC repeats larger than 23 units (17-20). We performed conditional logistic regression analysis on a subset of 2,287 screened cases and 4,162 controls including and excluding pathologic expansions carriers (n=144).

Estimation of the genetic variance tagged by all SNPs

We used GCTA (Genome-wide Complex Trait Analysis) software (29) to estimate the proportion of ALS phenotypic variance explained by autosomal SNPs distributed genome wide. Genetic relationships between pairs of individuals were calculated separately in the eight cohorts where all samples were previously tested for cryptic relatedness and excluded if the proportion of IBD (identical-by-descent) estimate was larger than 0.05 (Supplementary Table S2). We estimated the genetic relationship matrix (GRM) including imputed genotypes filtered by MAF > 0.01, posterior probability > 0.9 and information measure > 0.8. Filtered genotype data were submitted to GRM analysis and GRMs output data were used in the restricted maximum likelihood (REML) analysis. In each data set REML analysis was carried out by fitting the specific principal components (Supplementary Table S3) as covariates to control for possible population stratification. The variance estimate was transformed from the observed scale ($V(1)/V_p$) to a scale of liability ($V(1)/V_p_L$) by specifying a disease prevalence for ALS of 5 per 100,000 persons (1,2). The variance estimate of each cohort was combined by random effects meta-analysis weighted by sample size using the library *rmeta* in R toolset.

SUPPLEMENTARY MATERIAL

Supplementary Material is available at *HMG* online

ACKNOWLEDGEMENTS

Genetic and statistical analysis was performed using a Linux cluster supported by the National Institute for Health Research Biomedical Research Centre for Mental Health at the South London and Maudsley National Health Service Foundation Trust and Institute of Psychiatry King's College London. The SLAGEN Consortium was supported by funds from

the Istituto Superiore di Sanità [GW-SLA, Ricerca Finalizzata 2007 grant number 31]; I.F., C.G., B.C., J.F.P, V.S. were supported by the Istituto Superiore di Sanità [grant number 526D/34]; I.F. received salary support from “Amici del Centro Dino Ferrari”, V.S. acknowledge support from the “Amici del Centro Dino Ferrari”. A.R., C.G., B.C., N.T., V.S. were supported by the Istituto Superiore di Sanità [ALS-FTD, Ric; Finalizzata 2009 grant number 276]; G.Si. was supported by Monte dei Paschi di Siena Foundation [grant 39167] and the Istituto Superiore di Sanità [grant number RF-INP-2007-627227]; A.C. was supported in part by Federazione Italiana Giuoco Calcio and Fondazione Vialli e Mauro per la Sclerosi Laterale Amiotrofica onlus, Istituto Superiore di Sanità [grant number RF-PIE-2007-635695]; AF was supported by Progetto Regione Emilia Romagna-Università “RARER”. The Italian replica control samples were obtained from the “Parkinson Institute Biobank” (<http://www.parkinsonbiobank.com>) of the Telethon Genetic Biobank Network (<http://www.biobanknetwork.org>) supported by TELETHON Italy [grant number GTB07001] and by “Fondazione Grigioni per il Morbo di Parkinson”. ALSAGEN Consortium is supported by funds from the National Institutes for Health Research Biomedical Research Centre for Mental Health at the South London and Maudsley National Health Service Foundation Trust and Institute of Psychiatry, King's College London, the University of Leuven [grant number GOA/11/014], the Interuniversity Attraction Poles (IUAP) program P7/16 of the Belgian Federal Science Policy Office and the 7th Framework Programme of the European Union [ADAMS project, grant numbers HEALTH-F4-2009-242257 and FP7/2007-2013, grant agreements 259867 and 278611]. P.V.D. holds a clinical investigatorship from FWO-Vlaanderen; W.R. is supported through the E. von Behring Chair for Neuromuscular and Neurodegenerative Disorders. The UK National DNA Bank for Motor Neuron Disease Research was funded by the Motor Neurone Disease Association [grant 3/3], the Wellcome Trust [grant number 070122/A/02/Z] and the National Institute for Health Research

Dementias and Neurodegenerative Diseases Research Network. A.AC and J.F.P. receive salary support from the National Institute for Health Research Dementia Biomedical Research Unit at South London and Maudsley National Health Service Foundation Trust and King's College London. The RADIANT depression and bipolar case control studies (DeCC, BACC) were genotyped under Medical Research Council [grant number G0701420]. G.So. acknowledges EuroBiobank for ALS DNA samples; K.M. acknowledges Motor Neurone Disease Association UK, Queen Elizabeth Hospital Birmingham Charity. A.AC. thanks the Motor Neurone Disease Association of Great Britain and Northern Ireland, the ALS Association, the ALS Therapy Alliance and the Medical Research Council [grant number G0600974] for support. R.H.B. acknowledges support from the National Institute of Neurological Disorders and Stroke [grant number 5RO1-NS050557-05] and the National Institute of Neurological Disorders and Stroke American Recovery and Reinvestment Act Award [grant number RC2-NS070-342], the Angel Fund, the ALS Association, P2ALS, Project ALS, the Pierre L. de Bourgnecht ALS Research Foundation and the ALS Therapy Alliance. J.E.L. acknowledges grant support from National Institute of Health/ National Institute of Neurological Disorders and Stroke [grant number 1R01NS073873]. The Dutch, Belgian and Swedish GWAS data generation was supported by the Prinses Beatrix Fonds, VSB fonds, H. Kersten and M. Kersten (Kersten Foundation), The Netherlands ALS Foundation, and J.R. van Dijk and the Adessium Foundation. M.S. acknowledges support through EU via the Euro-MOTOR consortium and German BMBF Motoneuron disease network.

Conflict of interest statement

None declared.

Data Availability

Data was accessed by formal request to the consortia members. The complete summary meta-analysis data (SNP, genomic position, odds ratio and *P*-Value) is freely available at the ALSOD web site (<http://alsod.iop.kcl.ac.uk/>) in a searchable format (48).

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APPENDIX

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Legend to figures

Figure 1. Manhattan plot of the international ALS meta-analysis

Scatter plot of chromosome position (x axis) against $-\log_{10}$ GWAS meta-analysis P values (y axis) from imputed data. The threshold of genome-wide significance ($P = 5 \times 10^{-8}$) is indicated as a horizontal red line. At locus 9p21.2 18 SNPs close to *Corf72* gene lie above the red line (most significant SNP is rs3849943 with $P = 7.69 \times 10^{-9}$). Locus 17q11.2 shows SNP rs34517613 ($P = 1.11 \times 10^{-8}$) to be significantly associated. SNP rs1788776 at 18q11.2 is very close to the threshold of significance with $P = 7.67 \times 10^{-8}$. Manhattan plot was produced using *ggplot2* package in R.

Figure 2. Regional association plots of the associated loci

LD structure of the three regions associated with amyotrophic lateral sclerosis in the international GWAS meta-analysis. For each plot, the $-\log_{10} P$ values (left y axis) of SNPs are shown according to their chromosomal positions (x axis); the genetic recombination rates are shown on the right y axis. $-\log_{10} P$ values are shown for both genotyped and imputed SNPs distributed in a 0.8 megabase genomic region. The top SNP of each region is indicated as a diamond and SNPs colour reflects LD correlation (r^2). (a) locus at 9p21, (b) 17q11.2 and (c) 18q11.2. LD plots were generated by *LocusZoom v1.1*.

Figure 3. Heritability estimate across ALS-GWA studies

Heritability estimated in eight independent ALS cohorts. Forest plot shows heritability value and confidence interval (95%) calculated in each study at the reported prevalence of 5/100,000. Black boxes indicate the single studies and box sizes are proportional to the sample (N). Light blue bars specify lower and upper limit of 95% confidence interval.

TABLES

Table 1. Sample size and genotype platforms

Independent studies analyzed in the combined international meta-analysis were genotyped on different type of Illumina platforms. Italian data were collected by the SLAGEN Consortium (Italy) and the Dutch, Swedish, French, Belgian, Irish and North American, and Italian raw data were collected by the International ALS-GWAS Consortium (ALSAGEN). British cases were collected by UK National MND Bank samples and British controls by the RADIANT study (DeCC, BACCs) and NIH publicly available data from Coriell biobank (www.coriell.org).

Source	Sample Ancestry	Sample size (N)	Cases (N)	Control (N)	Genotyping Illumina Arrays	SNPs before imputation	SNPs after imputation
SLAGEN Consortium	Italian	3959	1982	1977	660K	657366	8391895
UMC Utrecht	Dutch	911	461	450	317K	317503	8328981
Utrecht, Umeå , Leuven	Dutch, Swedish, Belgian	2806	1364	1442	370K	317503	8372645
MGH, King's College , Evry	Northern American, British, French	3916	1710	2206	370K	307790	8370626
Beaumont Hospital, Dublin	Irish	432	221	211	550K	561466	8268635
NIH publicly available, Coriell biobank	Northern American	547	276	271	550K	555352	8325506
MND DNA bank and RADIANT study	British	2252	663	1589	610K	471994	8372429
NIH publicly available	Italian	747	500	247	550K	500002	8312642

Table 2. ALS-GWAS meta-analysis results

Summary of the most significant associated SNPs in the international ALS-GWAS meta-analysis from imputed genotypes. Meta-analysis (METAL) was performed weighting by the β -coefficient estimates and the inverse of their corresponding standard errors. Effect size of reference allele A1 is expressed with positive and negative symbols; +/- indicates the direction of the effect size (beta values from regression), where a plus symbol means that increasing frequency of allele A1 is correlated with increasing trait values and vice versa.

Chromosome	SNP	Position	A1	A2	APP	Info	MAF_ALS	MAF_CONT	OR	P-value	Direction	Low 95% CI	Up 95% CI
9	rs3849943	27533382	C	T	1	0.999	0.268	0.238	1.166	7.69E-09	+++++++	1.107	1.228
9	rs3849941	27528948	A	G	0.993	0.986	0.266	0.236	1.166	9.54E-09	++-----	1.106	1.229
9	rs2477523	27526894	T	A	0.997	0.992	0.269	0.239	1.163	1.06E-08	++-----	1.104	1.225
9	rs2453555	27553868	G	A	0.995	0.988	0.27	0.239	1.163	1.07E-08	++-----	1.104	1.225
17	rs34517613	23634379	C	T	0.922	0.734	0.108	0.129	0.822	1.11E-08	+++++++	0.769	0.879
17	rs12937920	23614784	C	T	0.951	0.854	0.125	0.149	0.83	1.24E-07	+++++++	0.775	0.889
17	rs9909055	23620255	C	T	1	1	0.217	0.249	0.857	1.98E-07	+++++++	0.809	0.908
17	rs16964203	23606837	A	C	0.978	0.954	0.224	0.257	0.86	2.55E-07	-----	0.812	0.911
17	rs2292633	23599507	G	A	0.956	0.906	0.221	0.254	0.86	2.82E-07	+++++++	0.812	0.911
17	rs35714695	23743915	G	A	0.973	0.934	0.169	0.192	0.845	3.21E-07	+++++++	0.792	0.901
17	rs739439	23747949	C	T	0.994	0.984	0.169	0.193	0.842	3.97E-07	+++++++	0.788	0.900
17	rs9900311	23619534	G	A	0.997	0.992	0.219	0.252	0.858	4.58E-07	+++++++	0.808	0.911
18	rs1788776	19498035	G	A	0.971	0.947	0.392	0.362	0.872	7.67E-08	+++++++	0.830	0.917
18	rs1629335	19475137	C	G	0.956	0.927	0.383	0.353	0.874	1.00E-07	-----	0.832	0.918
18	rs62093304	19489968	T	C	0.979	0.96	0.389	0.36	0.875	2.09E-07	-----+	0.832	0.920
18	rs12957850	19487958	G	T	0.988	0.976	0.408	0.382	0.878	2.64E-07	+++++++	0.836	0.923
18	rs1652373	19482645	G	T	1	1	0.409	0.382	0.877	2.66E-07	+++++++	0.834	0.922
18	rs12606211	19499478	C	T	0.979	0.961	0.408	0.381	0.88	3.29E-07	+++++++	0.838	0.924
18	rs1711468	19498537	C	A	0.97	0.945	0.41	0.384	0.882	4.16E-07	+++++++	0.840	0.926

Abbreviations

Amyotrophic lateral sclerosis (ALS), genome wide association studies (GWAS), Italian Consortium for the Genetics of ALS (SLAGEN), International Consortium on Amyotrophic Lateral Sclerosis Genetics (ALSGEN), single nucleotide polymorphisms (SNPs), odd ratio (OR), confidence interval (CI), minor allele frequency (MAF), linkage disequilibrium (LD), quality control (QC), identical by descent (IBD), observed quantiles versus the expected ones (Q-Q plot), principal components analysis (PCA), average posterior probability (APP), expression quantitative trait loci (eQTL), genome-wide complex trait analysis (GCTA), restricted maximum likelihood (REML), genetic relationship matrix (GRM).





